

Foundations of Machine Learning African Masters in Machine Intelligence

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Principal Component Analysis

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References

- Bishop: Pattern Recognition and Machine Learning, Chapter 12
- Deisenroth et al.: Mathematics for Machine Learning, Chapter 10 (https://mml-book.com)

Overview

Introduction

Setting

Maximum Variance Perspective

Projection Perspective

PCA Algorithm

PCA in High Dimensions

Probabilistic PCA

Related Models

High-Dimensional Data



- Real-world data is often high dimensional
- Challenges:
 - Difficult to analyze
 - Difficult to visualize
 - Difficult to interpret

Properties of High-dimensional Data



- Many dimensions are unnecessary
- Data often lives on a low-dimensional manifold
- ▶ Dimensionality reduction finds the relevant dimensions.

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Background: Coordinate Representations

Consider \mathbb{R}^2 with the canonical basis $e_1 = [1, 0]^\top$, $e_2 = [0, 1]^\top$.

$$x = \begin{bmatrix} 5 \\ 3 \end{bmatrix} = 5e_1 + 3e_2$$
 Linear combination of basis vectors

► **Coordinates** of *x* w.r.t. (*e*₁, *e*₂): [5,3]

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► **Coordinates** of *x* w.r.t. (*e*₁, *e*₂): [5,3]

Consider the vectors of the form

$$ilde{x} = \begin{bmatrix} 0 \\ z \end{bmatrix} \in \mathbb{R}^2, \quad z \in \mathbb{R}$$

Write them as $0e_1 + ze_2$.

- Only remember/store the coordinate/code *z* of the *e*₂ vector
 Compression
- Set of x vectors forms a vector subspace U ⊆ ℝ² with dim(U) = 1 because U = span[e₂].

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PCA Setting



- Dataset $\mathcal{X} := \{x_1, \ldots, x_N\}, x_n \in \mathbb{R}^D$
- ▶ Data matrix $X := [x_1, ..., x_N] \in \mathbb{R}^{D \times N}$ ▶ Often $N \times D$ matrix

PCA Setting



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- Without loss of generality: E[X] = 0 ▶ Centered data
 ▶ Data covariance matrix

PCA Setting



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- ▶ Data matrix $X := [x_1, ..., x_N] \in \mathbb{R}^{D \times N}$ ▶ Often $N \times D$ matrix
- Without loss of generality: E[X] = 0 ⇒ Centered data
 ⇒ Data covariance matrix S = ¹/_NXX^T ∈ ℝ^{D×D}
- Linear relationships between latent code *z* and data *x*:

$$z = B^{\top} x$$
, $\tilde{x} = B z$

•
$$\boldsymbol{B} = [\boldsymbol{b}_1, \dots, \boldsymbol{b}_M] \in \mathbb{R}^{D \times M}$$
 is an orthogonal matrix

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Low-Dimensional Embedding



- Find an *M*-dimensional subspace $U \subset \mathbb{R}^D$ onto which we project the data
- $\tilde{x} = \pi_U(x)$ is the projection of *x* onto *U*
- ▶ Find projections x̃ that are as similar to x as possible
 ▶ Find basis vectors b₁,..., b_M
- Compression loss incurs if *M* « *D*

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PCA Idea: Maximum Variance



 Project *D*-dimensional data *x* onto an *M*-dimensional subspace that retains as much information as possible
 Data compression

PCA Idea: Maximum Variance



- Project *D*-dimensional data *x* onto an *M*-dimensional subspace that retains as much information as possible
 Data compression
- Informally: information = diversity = variance
 Maximize variance in projected space (Hotelling 1933)

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PCA Objective: Maximum Variance

Linear relationships:

$$z = B^{\top} x$$
, $\tilde{x} = B z$

- $\boldsymbol{B} = [\boldsymbol{b}_1, \dots, \boldsymbol{b}_M] \in \mathbb{R}^{D \times M}$ is an orthogonal matrix
- Columns of **B** are an ONB of an *M*-dimensional subspace of \mathbb{R}^D

PCA Objective: Maximum Variance

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- $\boldsymbol{B} = [\boldsymbol{b}_1, \dots, \boldsymbol{b}_M] \in \mathbb{R}^{D \times M}$ is an orthogonal matrix
- Columns of **B** are an ONB of an *M*-dimensional subspace of \mathbb{R}^D
- ▶ Find B = [b₁,..., b_M] so that the variance in the projected space is maximized

$$\max_{\boldsymbol{b}_1,\dots,\boldsymbol{b}_M} \mathbb{V}[\boldsymbol{z}] = \max_{\boldsymbol{b}_1,\dots,\boldsymbol{b}_M} \mathbb{V}[\boldsymbol{B}^\top \boldsymbol{x}]$$

s.t. $\|\boldsymbol{b}_1\| = 1 = \dots = \|\boldsymbol{b}_M\|$

Constrained optimization problem

• Maximize variance of first coordinate of $z \in \mathbb{R}^M$:

$$V_1 := \mathbb{V}[z_1] = \frac{1}{N} \sum_{n=1}^N z_{n1}^2$$

▶ Empirical variance of the training dataset

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Empirical variance of the training dataset

• First coordinate of *z*^{*n*} is

$$z_{n1} = \boldsymbol{b}_1^\top \boldsymbol{x}_n$$

➤ Coordinate of orthogonal projection of *x_n* onto span[*b*₁]
 (1-dimensional subspace spanned by *b*₁)

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➤ Coordinate of orthogonal projection of *x_n* onto span[*b*₁]
 (1-dimensional subspace spanned by *b*₁)

$$\mathbb{V}[z_1] = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{b}_1^{\mathsf{T}} \mathbf{x}_n)^2 = \frac{1}{N} \sum_{n=1}^{N} \mathbf{b}_1^{\mathsf{T}} \mathbf{x}_n \mathbf{x}_n^{\mathsf{T}} \mathbf{b}_1$$
$$= \mathbf{b}_1^{\mathsf{T}} \left(\frac{1}{N} \sum_{n=1}^{N} \mathbf{x}_n \mathbf{x}_n^{\mathsf{T}} \right) \mathbf{b}_1 = \mathbf{b}_1^{\mathsf{T}} \mathbf{S} \mathbf{b}_1$$

Maximize variance

$$\max_{\bm{b}_1, \|\bm{b}_1\|^2 = 1} \mathbb{V}[z_1] = \max_{\bm{b}_1, \|\bm{b}_1\|^2 = 1} \bm{b}_1^\top \bm{S} \bm{b}_1$$

Maximize variance

$$\max_{\boldsymbol{b}_1, \|\boldsymbol{b}_1\|^2 = 1} \mathbb{V}[z_1] = \max_{\boldsymbol{b}_1, \|\boldsymbol{b}_1\|^2 = 1} \boldsymbol{b}_1^\top \boldsymbol{S} \boldsymbol{b}_1$$

• Lagrangian:

$$L(\boldsymbol{b}_1, \boldsymbol{\lambda}) = \boldsymbol{b}_1^\top \boldsymbol{S} \boldsymbol{b}_1 + \boldsymbol{\lambda}_1 (1 - \boldsymbol{b}_1^\top \boldsymbol{b}_1)$$

Discuss with your neighbors and find λ_1 and b_1

Maximize variance

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Discuss with your neighbors and find λ_1 and b_1

• Setting the gradients w.r.t. b_1 and λ_1 to **0** yields

$$egin{aligned} m{S}m{b}_1 &= \lambda_1m{b}_1 \ m{b}_1^{ op}m{b}_1 &= 1 \end{aligned}$$

- ► *b*¹ is an eigenvector of the data covariance matrix *S*
- λ₁ is the corresponding eigenvalue

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•
$$S\boldsymbol{b}_1 = \lambda_1 \boldsymbol{b}_1$$

 $\mathbb{V}[z_1] = \boldsymbol{b}_1^\top S \boldsymbol{b}_1 = \lambda_1 \boldsymbol{b}_1^\top \boldsymbol{b}_1 = \lambda_1$

▶ Variance retained by first coordinate corresponds to eigenvalue λ_1

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 \blacktriangleright Choose eigenvector b_1 associated with the largest eigenvalue

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- Projection:
- Coordinate:

Direction with Maximal Variance

Maximizing the variance means to choose the direction b_1 as the eigenvector of the data covariance matrix *S* that is associated with the largest eigenvalue λ_1 of *S*.

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▶ Variance retained by first coordinate corresponds to eigenvalue λ_1

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• Projection:
$$\tilde{x}_n = b_1 b_1^\top x_n$$

• Coordinate:
$$z_{n1} = \boldsymbol{b}_1^\top \boldsymbol{x}_n$$

Direction with Maximal Variance

Maximizing the variance means to choose the direction b_1 as the eigenvector of the data covariance matrix *S* that is associated with the largest eigenvalue λ_1 of *S*.

M-dimensional Subspace with Maximum Variance

General Result

The *M*-dimensional subspace of \mathbb{R}^D that retains the most variance is spanned by the *M* eigenvectors of the data covariance matrix *S* that are associated with the *M* largest eigenvalues of *S*. (e.g., Bishop 2006)

Example: MNIST Embedding (Training Set)



Embedding of handwritten '0' and '1' digits (28 × 28 pixels) into a two-dimensional subspace, spanned by the first two principal components.

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Example: MNIST Reconstruction (Test Set)



 Reconstructions of original digits as the number of principal components increases

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Refresher: Orthogonal Projection onto Subspaces

- Basis $\boldsymbol{b}_1, \ldots, \boldsymbol{b}_M$ of a subspace $U \subset \mathbb{R}^D$
- Define $\boldsymbol{B} = [\boldsymbol{b}_1, ..., \boldsymbol{b}_M] \in \mathbb{R}^{D \times M}$
- Project $x \in \mathbb{R}^D$ onto subspace U:

$$\pi_U(\mathbf{x}) = \tilde{\mathbf{x}} = \mathbf{B}(\mathbf{B}^\top \mathbf{B})^{-1} \mathbf{B}^\top \mathbf{x}$$

If *b*₁,..., *b*_M form an orthonormal basis (*b*^T_i *b*_j = δ_{ij}), then the projection simplifies to

$$\tilde{x} = BB^{\top}x$$

PCA Objective: Minimize Reconstruction Error



 Objective: Find orthogonal projection that minimizes the average squared projection/reconstruction error

$$J = \frac{1}{N} \sum_{n=1}^{N} \left\| \mathbf{x}_n - \tilde{\mathbf{x}}_n \right\|^2$$

where $\tilde{x}_n = \pi_U(x_n)$ is the projection of x_n onto U

Derivation (1)

• Assume an orthonormal basis of $\mathbb{R}^D = \operatorname{span}[\boldsymbol{b}_1, \dots, \boldsymbol{b}_D]$, such that $\boldsymbol{b}_i^\top \boldsymbol{b}_j = \delta_{ij}$

Derivation (1)

- Assume an orthonormal basis of $\mathbb{R}^D = \operatorname{span}[\boldsymbol{b}_1, \dots, \boldsymbol{b}_D]$, such that $\boldsymbol{b}_i^\top \boldsymbol{b}_j = \delta_{ij}$
- Every data point *x* can be written as a linear combination of the basis vectors:

$$\boldsymbol{x} = \sum_{d=1}^{D} \eta_d \boldsymbol{b}_d = \boldsymbol{B} \boldsymbol{\eta}, \quad \boldsymbol{B} = [\boldsymbol{b}_1, \dots, \boldsymbol{b}_D]$$
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▶ Rotation of the standard coordinates to a new coordinate system defined by the basis $(b_1, ..., b_D)$.

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▶ Original coordinates x_d are replaced by η_d , d = 1, ..., D

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▶ Rotation of the standard coordinates to a new coordinate system defined by the basis $(b_1, ..., b_D)$.

▶ Original coordinates x_d are replaced by η_d , d = 1, ..., D

• Obtain $\eta_d = \mathbf{x}^\top \mathbf{b}_d$, such that

$$oldsymbol{x} = \sum_{d=1}^{D} (oldsymbol{x}^{ op} oldsymbol{b}_d) oldsymbol{b}_d$$

Objective

Approximate

$$\boldsymbol{x} = \sum_{d=1}^{D} \eta_d \boldsymbol{b}_d$$
 with $\tilde{\boldsymbol{x}} = \sum_{m=1}^{M} z_m \boldsymbol{b}_m$

using *M* ≪ *D* many basis vectors → **Projection** onto a lower-dimensional subspace

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Derivation (3): Objective



Derivation (3): Objective



Choose coordinates *z_{mn}* and basis vectors *b*₁,..., *b_D* such that the average squared reconstruction error

$$J_M = \frac{1}{N} \sum_{n=1}^N \|\boldsymbol{x}_n - \tilde{\boldsymbol{x}}_n\|^2$$

is minimized

Derivation (3): Objective



Choose coordinates *z_{mn}* and basis vectors *b*₁,..., *b_D* such that the average squared reconstruction error

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is minimized

 \blacktriangleright Compute gradients of J_M w.r.t. all variables, set to **0**, solve

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Derivation (4): Optimal Coordinates

Necessary condition for optimum:

$$\frac{\partial J_M}{\partial z_{mn}} = 0 \implies z_{mn} = \boldsymbol{x}_n^\top \boldsymbol{b}_m, \qquad m = 1, \dots, M$$

- The optimal projection is the orthogonal projection
- The optimal coordinate *z_{mn}* is the orthogonal projection of *x_n* onto the one-dimensional subspace spanned by *b_m*
- (b₁,..., b_D) is ONB ▶ span[b_{M+1},..., b_D] is orthogonal complement of principal subspace (span[b₁,..., b_M])
- ► If

$$\boldsymbol{x}_n = \sum_{d=1}^D \eta_{dn} \boldsymbol{b}_d$$
 and $\tilde{\boldsymbol{x}}_n = \sum_{m=1}^M z_{mn} \boldsymbol{b}_m$

then $\eta_{mn} = z_{mn}$ for $m = 1, \ldots, M$

▶ Minimum error is given by the orthogonal projection of x_n onto the principal subspace spanned by b_1, \ldots, b_M

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Derivation (5): Displacement Vector



Approximation error only plays a role in dimensions M + 1, ..., D:

$$oldsymbol{x}_n - oldsymbol{ ilde{x}}_n = \sum_{j=M+1}^D ig(oldsymbol{x}_n^ opoldsymbol{b}_jig)oldsymbol{b}_j$$

Derivation (5): Displacement Vector



Approximation error only plays a role in dimensions M + 1, ..., D:

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▶ Displacement vector $x_n - \tilde{x}_n$ lies in orthogonal complement U^{\perp} of principal subspace U (linear combination of the b_j for j = M + 1, ..., D)

From the previous slide:

$$\boldsymbol{x}_n - \tilde{\boldsymbol{x}}_n = \sum_{j=M+1}^D (\boldsymbol{x}_n^\top \boldsymbol{b}_j) \boldsymbol{b}_j$$

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$$oldsymbol{x}_n - oldsymbol{ ilde{x}}_n = \sum_{j=M+1}^D (oldsymbol{x}_n^{ op} oldsymbol{b}_j) oldsymbol{b}_j$$

Let's compute our reconstruction error:

$$J_M = \frac{1}{N} \sum_{n=1}^N \|\boldsymbol{x}_n - \tilde{\boldsymbol{x}}_n\|^2 = \frac{1}{N} \sum_{n=1}^N (\boldsymbol{x}_n - \tilde{\boldsymbol{x}}_n)^\top (\boldsymbol{x}_n - \tilde{\boldsymbol{x}}_n)$$

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$$= \frac{1}{N} \sum_{n=1}^{N} \sum_{j=M+1}^{D} (\mathbf{x}_{n}^{\top} \mathbf{b}_{j})^{2}$$

From the previous slide:

$$\boldsymbol{x}_n - \tilde{\boldsymbol{x}}_n = \sum_{j=M+1}^D (\boldsymbol{x}_n^\top \boldsymbol{b}_j) \boldsymbol{b}_j$$

Let's compute our reconstruction error:

$$J_M = \frac{1}{N} \sum_{n=1}^N \|\mathbf{x}_n - \tilde{\mathbf{x}}_n\|^2 = \frac{1}{N} \sum_{n=1}^N (\mathbf{x}_n - \tilde{\mathbf{x}}_n)^\top (\mathbf{x}_n - \tilde{\mathbf{x}}_n)$$
$$= \frac{1}{N} \sum_{n=1}^N \sum_{j=M+1}^D (\mathbf{x}_n^\top \mathbf{b}_j)^2$$
$$= \sum_{j=M+1}^D \mathbf{b}_j^\top S \mathbf{b}_j$$

where $S = \frac{1}{N} \sum_{n=1}^{N} x_n x_n^{\top}$ is the data covariance matrix

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- What remains: Minimize J_M w.r.t. b_j under the constraint that the b_j form an orthonormal basis.
- Similar setting to maximum variance perspective: Instead of maximizing the variance in the principal subspace, we minimize the variance in the orthogonal complement of the principal subspace
- End up with **eigenvalue problem**:

$$Sb_j = \lambda_j b_j$$
, $j = D + 1, \dots, M$

▶ Find the eigenvectors *b*_{*i*} of the data covariance matrix *S*

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- Corresponding value of the squared reconstruction error:

$$J_M = \sum_{j=M+1}^D \lambda_j$$

i.e., the sum of the eigenvalues associated with eigenvectors not in the principle subspace

- ▶ Find the eigenvectors *b*_{*j*} of the data covariance matrix *S*
- Corresponding value of the squared reconstruction error:

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i.e., the sum of the eigenvalues associated with eigenvectors not in the principle subspace

Minimizing J_M requires us to choose the M eigenvectors as the principle subspace that are associated with the M largest eigenvalues.



• Objective: Project *x* onto an affine subspace μ + span[b_1].

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► Shift scenario to the origin (affine subspace ~→ vector subspace)

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• Shift *x* as well (onto $x - \mu$).



• Orthogonal projection of $x - \mu$ onto subspace spanned by b_1

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• Move projected point $\pi_{U_1}(x)$ back into original (affine) setting.

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Key Steps of PCA

- 1. Compute the empirical mean μ of the data
- 2. Mean subtraction: Replace all data points x_i with $\bar{x}_i = x_i \mu$.
- 3. Standardization: Divide the data by its standard deviation in each dimension: $\hat{X}^{(d)} = \bar{X}/\sigma(X^{(d)})$ for d = 1, ..., D.
- 4. Eigendecomposition of the data covariance matrix: Compute the eigenvectors (orthonormal) and eigenvalues of the data covariance matrix *S*
- 5. Orthogonal projection: Choose the eigenvectors associated with the M largest eigenvalues to be the basis of the principal subspace. Obtain \tilde{X}
- 6. Moving back to original data space: $\tilde{X}^{(d)} = \tilde{X}^{(d)}\sigma(X^{(d)}) + \mu_d$

PCA Algorithm



Dataset

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Mean subtraction



Standardization (variance 1 in each direction)



• Eigendecomposition of the data covariance matrix



Orthogonal projection onto the principal subspace



Moving back to the original data space

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PCA for High-Dimensional Data

- ▶ Fewer data points than dimensions, i.e., *N* < *D*.
- At least D N + 1 eigenvalues 0.
- Computation time for computing eigenvalues of data covariance matrix S: O(D³)
- ► Rephrase PCA

Reformulating PCA

▶ Define *X* to be the *D* × *N*-dimensional centered data matrix, whose *n*th row is $(x_n - \mathbb{E}[x])^\top$ ▶ Mean normalization

Reformulating PCA

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- Corresponding covariance: $S = \frac{1}{N} X X^{\top}$
Reformulating PCA

- ▶ Define *X* to be the *D* × *N*-dimensional centered data matrix, whose *n*th row is $(x_n \mathbb{E}[x])^\top$ ▶ Mean normalization
- Corresponding covariance: $S = \frac{1}{N} X X^{\top}$
- Corresponding eigenvector equation:

$$S\boldsymbol{b}_i = \lambda_i \boldsymbol{b}_i \iff \frac{1}{N} \boldsymbol{X} \boldsymbol{X}^\top \boldsymbol{b}_i = \lambda_i \boldsymbol{b}_i$$

Reformulating PCA

- ▶ Define *X* to be the *D* × *N*-dimensional centered data matrix, whose *n*th row is $(x_n \mathbb{E}[x])^\top$ ▶ Mean normalization
- Corresponding covariance: $S = \frac{1}{N}XX^{\top}$
- Corresponding eigenvector equation:

$$S\boldsymbol{b}_i = \lambda_i \boldsymbol{b}_i \iff \frac{1}{N} \boldsymbol{X} \boldsymbol{X}^\top \boldsymbol{b}_i = \lambda_i \boldsymbol{b}_i$$

• Transformation (left-multiply by X^{\top}):

$$\frac{1}{N} \boldsymbol{X} \boldsymbol{X}^{\mathsf{T}} \boldsymbol{b}_{i} = \lambda_{i} \boldsymbol{b}_{i} \quad \Longleftrightarrow \quad \frac{1}{N} \boldsymbol{X}^{\mathsf{T}} \boldsymbol{X} \underbrace{\boldsymbol{X}^{\mathsf{T}} \boldsymbol{b}_{i}}_{=:\boldsymbol{v}_{i}} = \lambda_{i} \underbrace{\boldsymbol{X}^{\mathsf{T}} \boldsymbol{b}_{i}}_{=:\boldsymbol{v}_{i}}$$

v_i is an eigenvector of the *N* × *N*-matrix ¹/_N*X*^T*X*, which has the same non-zero eigenvalues as the original covariance matrix.
 → Get eigenvalues in *O*(*N*³) instead of *O*(*D*³).

The new eigenvalue/eigenvector equation is

$$\frac{1}{N} \boldsymbol{X}^{\top} \boldsymbol{X} \boldsymbol{v}_i = \lambda_i \boldsymbol{v}_i$$

where $v_i = X^{\top} b_i$

The new eigenvalue/eigenvector equation is

$$\frac{1}{N} \boldsymbol{X}^{\top} \boldsymbol{X} \boldsymbol{v}_i = \lambda_i \boldsymbol{v}_i$$

where $v_i = X^{\top} b_i$

• We want to recover the original eigenvectors b_i of the data covariance matrix $S = \frac{1}{N}XX^{\top}$

The new eigenvalue/eigenvector equation is

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• Make sure to normalize Xv_i so that $||Xv_i|| = 1$

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Probabilistic PCA

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Latent Variable Perspective



Model:

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 , $oldsymbol{\epsilon} \sim \mathcal{N}ig(oldsymbol{0},\,\sigma^2oldsymbol{I}ig)$

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Generative process:

$$egin{aligned} & z \sim \mathcal{N}ig(\mathbf{0}, \, oldsymbol{I}ig) \ & x | z \sim \mathcal{N}ig(x | B z + \mu, \, \sigma^2 oldsymbol{I}ig) \end{aligned}$$

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- Deal with data dimensions that are missing at random by applying Bayes' theorem

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PPCA Likelihood (integrate out the latent variables):

$$p(\mathbf{x}|\mathbf{B},\boldsymbol{\mu},\sigma^2) = \int p(\mathbf{x}|\mathbf{z},\boldsymbol{\mu},\sigma^2)p(\mathbf{z})d\mathbf{z}$$
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▶ Is Gaussian with mean and covariance

$$\mathbb{E}[x] = \mathbb{E}_{z}[Bz + \mu + \epsilon] = \mu$$
$$\mathbb{V}[x] = \mathbb{V}_{z}[Bz + \mu + \epsilon] = BB^{\top} + \sigma^{2}I$$

Joint Distribution and Posterior

Joint distribution of observed and latent variables

$$p(\mathbf{x}, \mathbf{z} | \mathbf{B}, \boldsymbol{\mu}, \sigma^2) = \mathcal{N}\left(\begin{bmatrix} \mathbf{x} \\ \mathbf{z} \end{bmatrix} \middle| \begin{bmatrix} \boldsymbol{\mu} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \mathbf{B} \mathbf{B}^\top + \sigma^2 \mathbf{I} & \mathbf{B} \\ \mathbf{B}^\top & \mathbf{I} \end{bmatrix} \right)$$

Posterior via Gaussian conditioning:

$$p(\boldsymbol{z}|\boldsymbol{x}, \boldsymbol{B}, \boldsymbol{\mu}, \sigma^2) = \mathcal{N}(\boldsymbol{z} \mid \boldsymbol{m}, \boldsymbol{C})$$
$$\boldsymbol{m} = \boldsymbol{B}^\top (\boldsymbol{B}\boldsymbol{B}^\top + \sigma^2 \boldsymbol{I})^{-1} (\boldsymbol{x} - \boldsymbol{\mu})$$
$$\boldsymbol{C} = \boldsymbol{I} - \boldsymbol{B}^\top (\boldsymbol{B}\boldsymbol{B}^\top + \sigma^2 \boldsymbol{I})^{-1} \boldsymbol{B}$$

▶ For a new observation x_* compute the posterior on $p(z_*|x_*, X)$ and examine it (e.g., variance).

• Generate new (plausible) data from this posterior

Maximum Likelihood Estimation

- In PPCA, we can determine the parameters μ, B, σ² via maximum likelihood estimation. PPCA Likelihood: p(X|μ, B, σ²)
- Result (e.g., Tipping & Bishop (1999)):

$$\mu_{\rm ML} = \frac{1}{N} \sum_{n=1}^{N} x_n \quad \Longrightarrow \text{Sample mean}$$

$$B_{\rm ML} = T(\Lambda - \sigma^2 I)^{\frac{1}{2}} R$$

$$\sigma_{\rm ML}^2 = \frac{1}{D - M} \sum_{j=M+1}^{D} \lambda_j \quad \Longrightarrow \text{Average variance in orth. complement}$$

 For σ → 0 the maximum likelihood solution gives the same result as PCA (see mml-book.com)

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Related Models

- Factor analysis: Axis-aligned noise (instead of isotropic)
- ► Independent component analysis: Non-Gaussian prior p(z) = ∏_m p_m(z_m)
- Kernel PCA
- Bayesian PCA: Priors on parameters *B*, μ, σ²
 - Approximate
 inference



- Gaussian process latent variable model (GP-LVM): Replace linear mapping in Bayesian PCA with Gaussian process. Point estimate of z
- ▶ Bayesian GP-LVM maintains a distribution on *z* ▶ Approximate inference

Principal Component Analysis

Marc Deisenroth

@AIMS Rwanda, October 4, 2018





- ▶ PCA: Algorithm for linear dimensionality reduction
- Orthogonal projection of data onto a lower-dimensional subspace
 - Maximizes the variance of the projection
 - Minimizes the average squared projection/reconstruction error
- High-dimensional data
- Probabilistic PCA